

# Second Order Runge–Kutta Methods for Itô Stochastic Differential Equations

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## Abstract

A new class of stochastic Runge–Kutta methods for the weak approximation of the solution of Itô stochastic differential equation systems with a multi-dimensional Wiener process is introduced. As the main innovation, the number of stages of the methods does not depend on the dimension of the driving Wiener process and the number of the necessary random variables is reduced considerably. This reduces the computational effort significantly. Order conditions for the stochastic Runge–Kutta methods assuring weak convergence with order two are calculated by applying the colored rooted tree analysis due to the author. Further, some coefficients for explicit second order stochastic Runge–Kutta schemes are presented.

*Key words:* stochastic Runge–Kutta method, stochastic differential equation, multi-colored rooted tree analysis, weak approximation, numerical method

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## 1 Introduction

The development of derivative free approximation methods for solutions of stochastic differential equations (SDEs) is subject of recent research. For example, derivative free Runge–Kutta type schemes have been proposed for the strong approximation of solutions of SDEs in [1,2,5,9,11,16]. On the other hand, for the weak approximation of solutions of SDEs particular schemes have to be developed, see e.g. [5,6,9,17]. Recently, second order stochastic Runge–Kutta (SRK) methods for the weak approximation have been studied by e.g. Kloeden and Platen [5], Komori et. al. [7,8], Rößler [12–14] and Tocino and Vigo-Aguiar [19]. However, due to the knowledge of the author, up to now

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all proposed second order SRK methods suffer from an inefficiency if they are applied to SDEs with a multi-dimensional Wiener process because the number of stages, and thus the number of evaluations of the diffusion function per step, depends at least linearly on the dimension  $m$  of the driving Wiener process. This drawback becomes significant especially for high-dimensional problems. The aim of the present paper is to overcome this drawback by introducing a new class of efficient SRK methods. Essentially, these new SRK methods possess two advantages: Firstly, the number of stages and thus the number of evaluations of the drift and the diffusion functions per step is constant, i.e. independent of the dimension  $m$  of the driving Wiener process. Secondly, the number of random variables that have to be simulated is only  $2m - 1$  for each step. The paper is organized as follows: In Sections 2–4 we briefly review the main concept of the rooted tree analysis for the weak approximation introduced in [12,15]. Then, in Section 5, a new class of SRK methods is introduced and the rooted tree analysis provided in Sections 2–4 is applied to calculate order conditions. Further, some coefficients for explicit second order SRK schemes are presented. The paper closes with a numerical example in Section 6.

Let  $(\Omega, \mathcal{F}, P)$  be a probability space with a filtration  $(\mathcal{F}_t)_{t \geq 0}$  fulfilling the usual conditions and let  $\mathcal{I} = [t_0, T]$  for some  $0 \leq t_0 < T < \infty$ . We denote by  $(X_t)_{t \in \mathcal{I}}$  the solution of the  $d$ -dimensional Itô SDE system

$$X_t = X_{t_0} + \int_{t_0}^t a(s, X_s) ds + \sum_{j=1}^m \int_{t_0}^t b^j(s, X_s) dW_s^j \quad (1)$$

for  $d, m \geq 1$  with an  $m$ -dimensional Wiener process  $(W)_{t \geq 0}$ . Suppose  $X_{t_0} = x_0 \in \mathbb{R}^d$  is  $\mathcal{F}_{t_0}$ -measurable with  $E(\|X_{t_0}\|^{2l}) < \infty$  for some  $l \in \mathbb{N}$ . Throughout the paper we shall suppose that  $a, b^j : \mathcal{I} \times \mathbb{R}^d \rightarrow \mathbb{R}^d$  for  $j = 1, \dots, m$  are at least Lipschitz continuous and satisfy a linear growth condition w.r.t. the state variable  $x$ . Then the Existence and Uniqueness Theorem applies (see, e.g., [4]). In the following, let  $C_P^l(\mathbb{R}^d, \mathbb{R})$  denote the space of all  $g \in C^l(\mathbb{R}^d, \mathbb{R})$  with polynomial growth, i.e. there exists a constant  $C > 0$  and  $r \in \mathbb{N}$ , such that  $|\partial_x^i g(x)| \leq C(1 + \|x\|^{2r})$  for all  $x \in \mathbb{R}^d$  and any partial derivative of order  $i \leq l$  [5]. We say that  $g$  belongs to  $C_P^{k,l}(\mathcal{I} \times \mathbb{R}^d, \mathbb{R})$  if  $g \in C^{k,l}(\mathcal{I} \times \mathbb{R}^d, \mathbb{R})$  and  $g(t, \cdot) \in C_P^l(\mathbb{R}^d, \mathbb{R})$  holds uniformly in  $t \in \mathcal{I}$ . Let a discretization  $\mathcal{I}_h = \{t_0, t_1, \dots, t_N\}$  with  $t_0 < t_1 < \dots < t_N = T$  of the time interval  $\mathcal{I} = [t_0, T]$  with step sizes  $h_n = t_{n+1} - t_n$  for  $n = 0, 1, \dots, N - 1$  be given. Further, let  $h = \max_{0 \leq n < N} h_n$  denote the maximum step size.

**Definition 1.1** *An approximation process  $Y$  converges weakly with order  $p$  to  $X$  at time  $T$  as  $h \rightarrow 0$  if for each  $f \in C_P^{2(p+1)}(\mathbb{R}^d, \mathbb{R})$  exist a constant  $C_f$  and a finite  $\delta_0 > 0$  such that for each  $h \in ]0, \delta_0[$  holds:*

$$|E(f(X_T)) - E(f(Y(T)))| \leq C_f h^p. \quad (2)$$

## 2 Stochastic Runge–Kutta Methods

In [12], a very general class of stochastic Runge–Kutta methods has been introduced which covers the structure of well known schemes as well as of many future schemes: Let  $\mathcal{M}$  be a finite set of multi-indices with  $\kappa = |\mathcal{M}|$  elements and let  $\theta_\nu(h)$ ,  $\nu \in \mathcal{M}$ , be some random variables. Then, a general class of  $s$ -stage stochastic Runge–Kutta methods is given by  $Y_0 = x_0$  and

$$Y_{n+1} = Y_n + \sum_{i=1}^s z_i^{(0,0)} a \left( t_n + c_i^{(0,0)} h_n, H_i^{(0,0)} \right) \quad (3)$$

$$+ \sum_{i=1}^s \sum_{k=1}^m \sum_{\nu \in \mathcal{M}} z_i^{(k,\nu)} b^k \left( t_n + c_i^{(k,\nu)} h_n, H_i^{(k,\nu)} \right)$$

for  $n = 0, 1, \dots, N - 1$  with  $Y_n = Y(t_n)$ ,  $t_n \in \mathcal{I}_h$ , and

$$H_i^{(k,\nu)} = Y_n + \sum_{j=1}^s Z_{ij}^{(k,\nu),(0,0)} a \left( t_n + c_j^{(0,0)} h_n, H_j^{(0,0)} \right)$$

$$+ \sum_{j=1}^s \sum_{r=1}^m \sum_{\mu \in \mathcal{M}} Z_{ij}^{(k,\nu),(r,\mu)} b^r \left( t_n + c_j^{(r,\mu)} h_n, H_j^{(r,\mu)} \right)$$

for  $i = 1, \dots, s$ ,  $k = 0, 1, \dots, m$  and  $\nu \in \mathcal{M} \cup \{0\}$ . Here, let

$$z_i^{(0,0)} = \alpha_i h_n \quad z_i^{(k,\nu)} = \sum_{\iota \in \mathcal{M}} \gamma_i^{(\iota)(k,\nu)} \theta_\iota(h_n)$$

$$Z_{ij}^{(k,\nu),(0,0)} = A_{ij}^{(k,\nu),(0,0)} h_n \quad Z_{ij}^{(k,\nu),(r,\mu)} = \sum_{\iota \in \mathcal{M}} B_{ij}^{(\iota)(k,\nu),(r,\mu)} \theta_\iota(h_n)$$

for  $i, j = 1, \dots, s$  and let  $\alpha_i, \gamma_i^{(\iota)(k,\nu)}, A_{ij}^{(k,\nu),(0,0)}, B_{ij}^{(\iota)(k,\nu),(r,\mu)} \in \mathbb{R}$  be the coefficients of the SRK method. The weights can be defined by

$$c^{(k,\nu)} = A^{(k,\nu),(0,0)} e \quad (4)$$

with  $e = (1, \dots, 1)^T$ . If  $A_{ij}^{(k,\nu),(0,0)} = B_{ij}^{(\iota)(k,\nu),(r,\mu)} = 0$  for  $j \geq i$  then (3) is called an explicit SRK method, otherwise it is called implicit. We assume that the random variables  $\theta_\nu(h_n)$  satisfy the moment condition

$$\mathbb{E} \left( \theta_{\nu_1}^{p_1}(h_n) \cdot \dots \cdot \theta_{\nu_\kappa}^{p_\kappa}(h_n) \right) = O \left( h_n^{(p_1 + \dots + p_\kappa)/2} \right) \quad (5)$$

for all  $p_i \in \mathbb{N}_0$  and  $\nu_i \in \mathcal{M}$ ,  $1 \leq i \leq \kappa$ . The moment condition ensures a contribution of each random variable having an order of magnitude  $O(\sqrt{h})$ . This condition is in accordance with the order of magnitude of the increments of the Wiener process. Remark that in the case of  $b \equiv 0$ , the SRK method reduces to the well known deterministic Runge–Kutta method, so the introduced class of SRK methods turns out to be a generalization of deterministic Runge–Kutta methods.

### 3 Colored Rooted Tree Analysis

Next, we give a short outline of the colored rooted tree analysis following [12,15] which is applied in Section 5 for the calculation of order conditions for the SRK method (3). Without loss of generality, we restrict our considerations to autonomous  $d$ -dimensional Itô SDE systems with an  $m$ -dimensional Wiener process in this section. Let  $TS(\Delta)$  denote the set of colored rooted trees with a root of type  $\gamma = \otimes$  and which may additionally consist of some deterministic nodes of type  $\tau = \bullet$  and stochastic nodes of type  $\sigma_{j_k} = \circ_{j_k}$  with a variable index  $j_k \in \{1, \dots, m\}$ . The variable index  $j_k$  is associated with the  $j_k$ th component of the  $m$ -dimensional driving Wiener process of the considered SDE. If not stated otherwise, each stochastic node has its own variable index. So, if we have a tree with  $s$  stochastic nodes then to each stochastic node corresponds exactly one of the indices  $j_1, \dots, j_s$ .

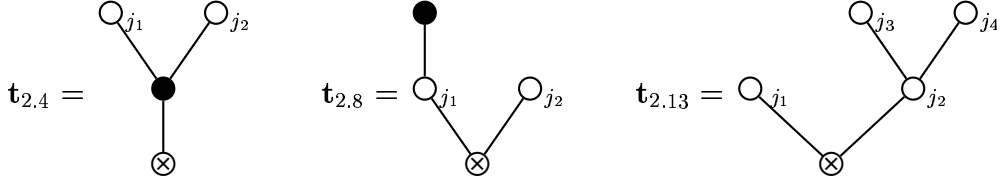


Fig. 1. Some elements of  $TS(\Delta)$  with  $j_1, j_2, j_3, j_4 \in \{1, \dots, m\}$ .

Every tree  $\mathbf{t} \in TS(\Delta)$  can be written by a combination of three different brackets: If  $\mathbf{t}_1, \dots, \mathbf{t}_k$  are colored trees then we denote by  $(\mathbf{t}_1, \dots, \mathbf{t}_k)$ ,  $[\mathbf{t}_1, \dots, \mathbf{t}_k]$  and  $\{\mathbf{t}_1, \dots, \mathbf{t}_k\}_j$  the tree in which  $\mathbf{t}_1, \dots, \mathbf{t}_k$  are each joined by a single branch to  $\otimes$ ,  $\bullet$  and  $\circ_j$ , respectively. Here, the order of the subtrees  $\mathbf{t}_1, \dots, \mathbf{t}_k$  does not matter since any order leads to equivalent trees. Therefore, we obtain  $\mathbf{t}_{2.4} = ([\circ_{j_1}, \circ_{j_2}]) = ([\sigma_{j_1}, \sigma_{j_2}])$ ,  $\mathbf{t}_{2.8} = (\{\bullet\}_{j_1}, \circ_{j_2}) = (\{\tau\}_{j_1}, \sigma_{j_2})$  and  $\mathbf{t}_{2.13} = (\circ_{j_1}, \{\circ_{j_3}, \circ_{j_4}\}_{j_2}) = (\sigma_{j_1}, \{\sigma_{j_3}, \sigma_{j_4}\}_{j_2})$  for the trees in Figure 1.

In the following, let  $l(\mathbf{t})$  be the number of nodes of  $\mathbf{t} \in TS(\Delta)$ . Then, we denote by  $d(\mathbf{t})$  the number of deterministic nodes, by  $s(\mathbf{t})$  the number of stochastic nodes of  $\mathbf{t} \in TS(\Delta)$  and it holds  $l(\mathbf{t}) = d(\mathbf{t}) + s(\mathbf{t}) + 1$ . The order  $\rho(\mathbf{t})$  of the tree  $\mathbf{t} \in TS(\Delta)$  is defined as  $\rho(\mathbf{t}) = d(\mathbf{t}) + \frac{1}{2}s(\mathbf{t})$  with  $\rho(\gamma) = 0$ . For example, it holds  $\rho(\mathbf{t}_{2.4}) = \rho(\mathbf{t}_{2.8}) = \rho(\mathbf{t}_{2.13}) = 2$ .

Now, let  $LTS(\Delta)$  denote the set of monotonically labelled trees, i.e. where the nodes are monotonically numbered starting with number one at the root of the tree. Then,  $\alpha_\Delta(\mathbf{t})$  is the cardinality of  $\mathbf{t}$ , i.e. the number of possibilities of monotonically labelling the nodes of  $\mathbf{t}$  with numbers  $1, \dots, l(\mathbf{t})$ . For example, for  $\mathbf{t}_{1.2} = (\sigma_{j_1}, \sigma_{j_2})$  exists only one possible monotonically labelling  $(\sigma_{j_1}^2, \sigma_{j_2}^3)^1$  and thus  $\alpha_\Delta(\mathbf{t}_{1.2}) = 1$ . In contrast to this, for  $\mathbf{t}_{1.5.3} = (\tau, \sigma_{j_1})$  holds  $\alpha_\Delta(\mathbf{t}_{1.5.3}) = 2$ . Although  $(\tau^2, \sigma_{j_1}^3)^1$  and  $(\sigma_{j_1}^3, \tau^2)^1$  are equivalent trees, there exist two different labelled trees  $(\tau^2, \sigma_{j_1}^3)^1$  and  $(\tau^3, \sigma_{j_1}^2)^1$ . So one has to dis-

tinguish between the labels of deterministic and stochastic nodes (see [12] for details). Further, it holds  $\alpha_\Delta(\mathbf{t}_{2.4}) = 1$ ,  $\alpha_\Delta(\mathbf{t}_{2.8}) = 3$  and  $\alpha_\Delta(\mathbf{t}_{2.13}) = 4$ .

Now, we assign to each tree  $\mathbf{t} \in TS(\Delta)$  an elementary differential which is defined recursively by  $F(\gamma)(x) = f(x)$ ,  $F(\tau)(x) = a(x)$ ,  $F(\sigma_j)(x) = b^j(x)$  and

$$F(\mathbf{t})(x) = \begin{cases} f^{(k)}(x) \cdot (F(\mathbf{t}_1)(x), \dots, F(\mathbf{t}_k)(x)) & \text{for } \mathbf{t} = (\mathbf{t}_1, \dots, \mathbf{t}_k) \\ a^{(k)}(x) \cdot (F(\mathbf{t}_1)(x), \dots, F(\mathbf{t}_k)(x)) & \text{for } \mathbf{t} = [\mathbf{t}_1, \dots, \mathbf{t}_k] \\ b^{j^{(k)}}(x) \cdot (F(\mathbf{t}_1)(x), \dots, F(\mathbf{t}_k)(x)) & \text{for } \mathbf{t} = \{\mathbf{t}_1, \dots, \mathbf{t}_k\}_j \end{cases} \quad (6)$$

Here  $f^{(k)}$ ,  $a^{(k)}$  and  $b^{j^{(k)}}$  define a symmetric  $k$ -linear differential operator, and one can choose the sequence of subtrees  $\mathbf{t}_1, \dots, \mathbf{t}_k$  in an arbitrary order. For example, the  $I$ th component of  $a^{(k)} \cdot (F(\mathbf{t}_1), \dots, F(\mathbf{t}_k))$  can be written as

$$(a^{(k)} \cdot (F(\mathbf{t}_1), \dots, F(\mathbf{t}_k)))^I = \sum_{J_1, \dots, J_k=1}^d \frac{\partial^k a^I}{\partial x^{J_1} \dots \partial x^{J_k}} (F^{J_1}(\mathbf{t}_1), \dots, F^{J_k}(\mathbf{t}_k))$$

where the components of vectors are denoted by superscript indices, which are chosen as capitals. For example, we obtain for  $\mathbf{t}_{2.8}$  the elementary differential

$$F(\mathbf{t}_{2.8}) = f''(b^{j_1'}(a), b^{j_2}) = \sum_{J_1, J_2=1}^d \frac{\partial^2 f}{\partial x^{J_1} \partial x^{J_2}} \left( \sum_{K_1=1}^d \frac{\partial b^{J_1, j_1}}{\partial x^{K_1}} a^{K_1} \cdot b^{J_2, j_2} \right).$$

**Definition 3.1** Let  $TS(I)$  denote the set of trees  $\mathbf{t} \in TS(\Delta)$  with a root of type  $\gamma$  which can be build by finite many steps of the form

- a) adding a deterministic node of type  $\tau$ , or
- b) adding two stochastic nodes of type  $\sigma_{j_k}$ , both with the same new variable index  $j_k$  for some  $k \in \mathbb{N}$ , whereas neither of the two nodes is allowed to be directly connected by an edge with the other one.

Let  $LTS(I)$  denote the set of labelled trees  $\mathbf{t} \in TS(I)$  with the nodes labelled in the same order as they are added. Then,  $\alpha_I(\mathbf{t})$  is the number of all possible different monotonically labels of  $\mathbf{t} \in TS(I)$  with  $\alpha_I(\mathbf{t}) = 0$  if  $\mathbf{t} \notin TS(I)$ .

For example,  $(\{\sigma_{j_1}^3\}_{j_1}^2, \{\sigma_{j_2}^5\}_{j_2}^4)^1 \notin LTS(I)$  while  $(\{\sigma_{j_2}^5\}_{j_2}^2, \{\sigma_{j_1}^4\}_{j_1}^3)^1 \in LTS(I)$ . The following result holds due to Theorem 3.2 and Proposition 5.1 in [15].

**Theorem 3.2** For  $p \in \mathbb{N}_0$ ,  $f, a^i, b^{i,j} \in C_P^{2(p+1)}(\mathbb{R}^d, \mathbb{R})$ ,  $i = 1, \dots, d$ ,  $j = 1, \dots, m$ , and for  $t \in [t_0, T]$  with  $h = t - t_0$  the following truncated expansion holds:

$$E^{t_0, x_0}(f(X_t)) = \sum_{\substack{\mathbf{t} \in TS(I) \\ \rho(\mathbf{t}) \leq p}} \sum_{j_1, \dots, j_{s(\mathbf{t})/2}=1}^m \frac{\alpha_I(\mathbf{t}) F(\mathbf{t})(x_0)}{2^{s(\mathbf{t})/2} \rho(\mathbf{t})!} h^{\rho(\mathbf{t})} + \mathcal{O}(h^{p+1}). \quad (7)$$

Next, we give an expansion for the approximation process  $(Y(t))_{t \in \mathcal{I}_h}$  defined by the SRK method (3). For  $\mathbf{t} \in TS(\Delta)$  let the density  $\gamma(\mathbf{t})$  be defined recursively by  $\gamma(\mathbf{t}) = 1$  if  $l(\mathbf{t}) = 1$  and

$$\gamma(\mathbf{t}) = \begin{cases} \prod_{i=1}^{\lambda} \gamma(\mathbf{t}_i) & \text{if } \mathbf{t} = (\mathbf{t}_1, \dots, \mathbf{t}_\lambda), \\ l(\mathbf{t}) \prod_{i=1}^{\lambda} \gamma(\mathbf{t}_i) & \text{if } \mathbf{t} = [\mathbf{t}_1, \dots, \mathbf{t}_\lambda] \text{ or } \mathbf{t} = \{\mathbf{t}_1, \dots, \mathbf{t}_\lambda\}_j. \end{cases}$$

Since the expansion for  $(Y(t))_{t \in \mathcal{I}_h}$  contains the coefficients and the random variables of the SRK method, we define a coefficient function  $\Phi_S$  which assigns to every tree  $\mathbf{t} \in TS(\Delta)$  an elementary weight:

$$\Phi_S(\mathbf{t}) = \begin{cases} \prod_{i=1}^{\lambda} \Phi_S(\mathbf{t}_i) & \text{if } \mathbf{t} = (\mathbf{t}_1, \dots, \mathbf{t}_\lambda) \\ z^{(0,0)T} \prod_{i=1}^{\lambda} \Psi^{(0,0)}(\mathbf{t}_i) & \text{if } \mathbf{t} = [\mathbf{t}_1, \dots, \mathbf{t}_\lambda] \\ \sum_{\nu \in \mathcal{M}} z^{(k,\nu)T} \prod_{i=1}^{\lambda} \Psi^{(k,\nu)}(\mathbf{t}_i) & \text{if } \mathbf{t} = \{\mathbf{t}_1, \dots, \mathbf{t}_\lambda\}_k \end{cases} \quad (8)$$

where  $\Phi_S(\gamma) = 1$ ,  $\Psi^{(k,\nu)}(\emptyset) = e$  with  $\tau = [\emptyset]$ ,  $\sigma_k = \{\emptyset\}_k$  and

$$\Psi^{(k,\nu)}(\mathbf{t}) = \begin{cases} Z^{(k,\nu),(0,0)} \prod_{i=1}^{\lambda} \Psi^{(0,0)}(\mathbf{t}_i) & \text{if } \mathbf{t} = [\mathbf{t}_1, \dots, \mathbf{t}_\lambda] \\ \sum_{\mu \in \mathcal{M}} Z^{(k,\nu),(r,\mu)} \prod_{i=1}^{\lambda} \Psi^{(r,\mu)}(\mathbf{t}_i) & \text{if } \mathbf{t} = \{\mathbf{t}_1, \dots, \mathbf{t}_\lambda\}_r \end{cases}. \quad (9)$$

Here  $e = (1, \dots, 1)^T$  and the product of vectors in (9) is defined by component-wise multiplication, i.e.  $(a_1, \dots, a_n) \cdot (b_1, \dots, b_n) = (a_1 b_1, \dots, a_n b_n)$ . Remark that  $TS(I) \subset TS(\Delta)$ . Further, each tree  $\mathbf{t} \in TS(\Delta)$  has  $s(\mathbf{t})$  different variable indices  $j_1, \dots, j_{s(\mathbf{t})}$  while a tree  $\mathbf{u} \in TS(I)$  has only  $s(\mathbf{u})/2$  different variable indices. Then Proposition 6.1 in [12] yields:

**Proposition 3.3** *Let  $(Y(t))_{t \in \mathcal{I}_h}$  be defined by the SRK method (3). Assume that for the random variables holds  $\theta_\iota(h) = \sqrt{h} \cdot \vartheta_\iota$  for  $\iota \in \mathcal{M}$  with some bounded random variables  $\vartheta_\iota$ . Then for  $p \in \mathbb{N}_0$ ,  $f, a^i, b^{i,j} \in C_P^{2(p+1)}(\mathbb{R}^d, \mathbb{R})$  for  $i = 1, \dots, d$ ,  $j = 1, \dots, m$  and for  $t \in [t_0, T]$  with  $h = t - t_0$  holds:*

$$\mathbb{E}^{t_0, x_0} (f(Y(t))) = \sum_{\substack{\mathbf{t} \in TS(\Delta) \\ \rho(\mathbf{t}) \leq p + \frac{1}{2}}} \sum_{j_1, \dots, j_{s(\mathbf{t})}=1}^m \frac{\alpha_\Delta(\mathbf{t}) \gamma(\mathbf{t}) F(\mathbf{t})(x_0) \mathbb{E}(\Phi_S(\mathbf{t}))}{(l(\mathbf{t}) - 1)!} + \mathcal{O}(h^{p+1}).$$

## 4 Order Conditions for Stochastic Runge–Kutta Methods

Now, we apply the rooted tree expansions of the solution and the approximation processes in order to yield order conditions for the SRK method (3).

**Definition 4.1** Let  $|\mathbf{t}|$  denote the tree which is obtained if the nodes  $\sigma_{j_i}$  of  $\mathbf{t}$  are replaced by  $\sigma$ , i.e. by omitting all variable indices. Let a tree  $\mathbf{t} \in TS(I)$  with variable indices  $j_1, \dots, j_{s(\mathbf{t})/2}$  be given and let  $\mathbf{u} \in TS(\Delta)$  with variable indices  $\hat{j}_1, \dots, \hat{j}_{s(\mathbf{u})}$  denote the tree which is equivalent to  $\mathbf{t}$  except for the variable indices, i.e.  $|\mathbf{t}| \sim |\mathbf{u}|$  with  $s(\mathbf{t}) = s(\mathbf{u})$ . For a fixed choice of correlations of type  $j_k = j_l$  or  $j_k \neq j_l$ ,  $1 \leq k < l \leq s(\mathbf{t})/2$ , between the indices  $j_1, \dots, j_{s(\mathbf{t})/2}$ , let  $\beta(\mathbf{t})$  denote the number of all possible correlations between the indices  $\hat{j}_1, \dots, \hat{j}_{s(\mathbf{u})}$  of tree  $\mathbf{u}$  such that  $\mathbf{t} \sim \mathbf{u}$  holds. In the case of  $s(\mathbf{t}) = 0$  or  $\mathbf{t} \in TS(\Delta) \setminus TS(I)$  define  $\beta(\mathbf{t}) = 1$ .

Note that in case of  $m = 1$  we have  $\beta(\mathbf{t}) = 1$  for all  $\mathbf{t} \in TS(I)$ . For example, for  $\mathbf{t} = (\sigma_{j_1}, \sigma_{j_2}, \{\sigma_{j_2}\}_{j_1}) \in TS(I)$  and  $\mathbf{u} = (\sigma_{\hat{j}_1}, \sigma_{\hat{j}_2}, \{\sigma_{\hat{j}_2}\}_{\hat{j}_1}) \in TS(\Delta)$ , two cases have to be considered. On the one hand we have the correlation  $j_1 = j_2$  for  $\mathbf{t}$  where we get the only possible correlation  $\hat{j}_1 = \hat{j}_2 = \hat{j}_3 = \hat{j}_4$  for  $\mathbf{u}$ , i.e.  $\beta(\mathbf{t}) = 1$ . On the other hand we have  $j_1 \neq j_2$  as a correlation for  $\mathbf{t}$  allowing us two different correlations  $\hat{j}_1 = \hat{j}_3 \neq \hat{j}_2 = \hat{j}_4$  and  $\hat{j}_2 = \hat{j}_3 \neq \hat{j}_1 = \hat{j}_4$  for  $\mathbf{u}$ . Thus we get  $\beta(\mathbf{t}) = 2$  in the latter case.

The following theorem yields conditions for the coefficients and the random variables of the SRK method (3) such that convergence with some order  $p$  in the weak sense is assured (see Theorem 6.4 in [12]).

**Theorem 4.2** Let  $p \in \mathbb{N}$ ,  $a^i, b^{i,j} \in C_P^{p+1, 2p+2}(\mathcal{I} \times \mathbb{R}^d, \mathbb{R})$  for  $i = 1, \dots, d$ ,  $j = 1, \dots, m$ . Then the SRK method (3) with step size  $h$  is of weak order  $p$ , if for all  $\mathbf{t} \in TS(\Delta)$  with  $\rho(\mathbf{t}) \leq p + \frac{1}{2}$  and all correlations of type  $j_k = j_l$  or  $j_k \neq j_l$ ,  $1 \leq k < l \leq s(\mathbf{t})$ , between the indices  $j_1, \dots, j_{s(\mathbf{t})} \in \{1, \dots, m\}$  of  $\mathbf{t}$  holds

$$\mathbb{E}(\Phi_S(\mathbf{t})) = \frac{\alpha_I(\mathbf{t}) \cdot (l(\mathbf{t}) - 1)! \cdot h^{\rho(\mathbf{t})}}{\alpha_\Delta(\mathbf{t}) \cdot \beta(\mathbf{t}) \cdot \gamma(\mathbf{t}) \cdot 2^{s(\mathbf{t})/2} \cdot \rho(\mathbf{t})!} \quad (10)$$

provided (4) and (5) hold and if the approximation  $Y$  has uniformly bounded moments w.r.t. the number  $N$  of steps.

**Remark 4.3** The approximation  $Y$  by the SRK method (3) has uniformly bounded moments if bounded random variables are used by the method, if (5) is fulfilled and if  $\mathbb{E}(z^{(k,\nu)T} e) = 0$  holds for  $1 \leq k \leq m$  and  $\nu \in \mathcal{M}$  (see [12]). Further, Theorem 4.2 provides uniform weak convergence with order  $p$  in the case of a non-random time discretization  $\mathcal{I}_h$  [12].

## 5 Order Two Stochastic Runge–Kutta Methods for Itô SDEs

A second order SRK method for the weak approximation of the solution of the Itô SDE (1) is analyzed in this section. Therefore, we introduce a new class of SRK methods where the number of stages is independent of the dimension  $m$  of the driving Wiener process. We define the  $d$ -dimensional approximation process  $Y$  with  $Y_n = Y(t_n)$  for  $t_n \in \mathcal{I}_h$  by the following SRK method of  $s$  stages with  $Y_0 = x_0$  and

$$\begin{aligned}
Y_{n+1} = Y_n &+ \sum_{i=1}^s \alpha_i a(t_n + c_i^{(0)} h_n, H_i^{(0)}) h_n \\
&+ \sum_{i=1}^s \sum_{k=1}^m \beta_i^{(1)} b^k(t_n + c_i^{(1)} h_n, H_i^{(k)}) \hat{I}_{(k)} \\
&+ \sum_{i=1}^s \sum_{k=1}^m \beta_i^{(2)} b^k(t_n + c_i^{(1)} h_n, H_i^{(k)}) \frac{\hat{I}_{(k,k)}}{\sqrt{h_n}} \\
&+ \sum_{i=1}^s \sum_{k=1}^m \beta_i^{(3)} b^k(t_n + c_i^{(2)} h_n, \hat{H}_i^{(k)}) \hat{I}_{(k)} \\
&+ \sum_{i=1}^s \sum_{k=1}^m \beta_i^{(4)} b^k(t_n + c_i^{(2)} h_n, \hat{H}_i^{(k)}) \sqrt{h_n}
\end{aligned} \tag{11}$$

for  $n = 0, 1, \dots, N - 1$  with stage values

$$\begin{aligned}
H_i^{(0)} &= Y_n + \sum_{j=1}^s A_{ij}^{(0)} a(t_n + c_j^{(0)} h_n, H_j^{(0)}) h_n \\
&+ \sum_{j=1}^s \sum_{l=1}^m B_{ij}^{(0)} b^l(t_n + c_j^{(1)} h_n, H_j^{(l)}) \hat{I}_{(l)} \\
H_i^{(k)} &= Y_n + \sum_{j=1}^s A_{ij}^{(1)} a(t_n + c_j^{(0)} h_n, H_j^{(0)}) h_n \\
&+ \sum_{j=1}^s B_{ij}^{(1)} b^k(t_n + c_j^{(1)} h_n, H_j^{(k)}) \sqrt{h_n} \\
\hat{H}_i^{(k)} &= Y_n + \sum_{j=1}^s A_{ij}^{(2)} a(t_n + c_j^{(0)} h_n, H_j^{(0)}) h_n \\
&+ \sum_{j=1}^s \sum_{\substack{l=1 \\ l \neq k}}^m B_{ij}^{(2)} b^l(t_n + c_j^{(1)} h_n, H_j^{(l)}) \frac{\hat{I}_{(k,l)}}{\sqrt{h_n}}
\end{aligned}$$

for  $i = 1, \dots, s$  and  $k = 1, \dots, m$ . The random variables are defined by

$$\hat{I}_{(k,l)} = \begin{cases} \frac{1}{2}(\hat{I}_{(k)}\hat{I}_{(l)} - \sqrt{h_n}\tilde{I}_{(k)}) & \text{if } k < l \\ \frac{1}{2}(\hat{I}_{(k)}\hat{I}_{(l)} + \sqrt{h_n}\tilde{I}_{(l)}) & \text{if } l < k \\ \frac{1}{2}(\hat{I}_{(k)}^2 - h_n) & \text{if } k = l \end{cases} \tag{12}$$



for  $1 \leq k, l \leq m$  with independent random variables  $\hat{I}_{(k)}$ ,  $1 \leq k \leq m$ , such that

$$\mathbb{E}(\hat{I}_{(k)}^q) = \begin{cases} 0 & \text{for } q \in \{1, 3, 5\} \\ (q-1)h^{q/2} & \text{for } q \in \{2, 4\} \\ \mathcal{O}(h^{q/2}) & \text{for } q \geq 6 \end{cases} \quad (13)$$

and random variables  $\tilde{I}_{(k)}$ ,  $1 \leq k \leq m-1$ , possessing the moments

$$\mathbb{E}(\tilde{I}_{(k)}^q) = \begin{cases} 0 & \text{for } q \in \{1, 3\} \\ h & \text{for } q = 2 \\ \mathcal{O}(h^{q/2}) & \text{for } q \geq 4 \end{cases}. \quad (14)$$

Thus, only  $2m-1$  independent random variables are needed for each step. For example, we can choose  $\hat{I}_{(k)}$  as three point distributed random variables with  $\mathbb{P}(\hat{I}_{(k)} = \pm\sqrt{3h_n}) = \frac{1}{6}$  and  $\mathbb{P}(\hat{I}_{(k)} = 0) = \frac{2}{3}$ . The random variables  $\tilde{I}_{(k)}$  can be defined by a two point distribution with  $\mathbb{P}(\tilde{I}_{(k)} = \pm\sqrt{h}) = \frac{1}{2}$ .

The coefficients of the SRK method (11) can be represented by an extended Butcher array taking the form:

$c^{(0)}$	$A^{(0)}$	$B^{(0)}$	
$c^{(1)}$	$A^{(1)}$	$B^{(1)}$	
$c^{(2)}$	$A^{(2)}$	$B^{(2)}$	
	$\alpha^T$	$\beta^{(1)T}$	$\beta^{(2)T}$
		$\beta^{(3)T}$	$\beta^{(4)T}$

Applying the rooted tree analysis presented in section 3 and section 4, we can calculate order conditions for the SRK method (11).

**Theorem 5.1** *Let  $a^i, b^{i,j} \in C_P^{2,4}(\mathcal{I} \times \mathbb{R}^d, \mathbb{R})$  for  $1 \leq i \leq d, 1 \leq j \leq m$ . If the coefficients of the stochastic Runge–Kutta method (11) fulfill the equations*

1.  $\alpha^T e = 1$
2.  $\beta^{(4)T} e = 0$
3.  $\beta^{(3)T} e = 0$
4.  $(\beta^{(1)T} e)^2 = 1$
5.  $\beta^{(2)T} e = 0$
6.  $\beta^{(1)T} B^{(1)} e = 0$
7.  $\beta^{(4)T} A^{(2)} e = 0$
8.  $\beta^{(3)T} B^{(2)} e = 0$
9.  $\beta^{(4)T} (B^{(2)} e)^2 = 0$

*then the method attains order 1.0 for the weak approximation of the solution of the Itô SDE (1). Further, if  $a^i, b^{i,j} \in C_P^{3,6}(\mathcal{I} \times \mathbb{R}^d, \mathbb{R})$  for  $1 \leq i \leq d, 1 \leq j \leq m$  and if in addition the equations*

10.  $\alpha^T A^{(0)} e = \frac{1}{2}$
11.  $\alpha^T (B^{(0)} e)^2 = \frac{1}{2}$
12.  $(\beta^{(1)T} e)(\alpha^T B^{(0)} e) = \frac{1}{2}$
13.  $(\beta^{(1)T} e)(\beta^{(1)T} A^{(1)} e) = \frac{1}{2}$

$$\begin{array}{ll}
14. & \beta^{(3)T} A^{(2)} e = 0 \\
15. & \beta^{(2)T} B^{(1)} e = 1 \\
16. & \beta^{(4)T} B^{(2)} e = 1 \\
17. & (\beta^{(1)T} e)(\beta^{(1)T} (B^{(1)} e)^2) = \frac{1}{2} \\
18. & (\beta^{(1)T} e)(\beta^{(3)T} (B^{(2)} e)^2) = \frac{1}{2} \\
19. & \beta^{(1)T} (B^{(1)} (B^{(1)} e)) = 0 \\
20. & \beta^{(3)T} (B^{(2)} (B^{(1)} e)) = 0 \\
21. & \beta^{(3)T} (B^{(2)} (B^{(1)} (B^{(1)} e))) = 0 \\
22. & \beta^{(1)T} (A^{(1)} (B^{(0)} e)) = 0 \\
23. & \beta^{(3)T} (A^{(2)} (B^{(0)} e)) = 0 \\
24. & \beta^{(4)T} (A^{(2)} e)^2 = 0 \\
25. & \beta^{(4)T} (A^{(2)} (A^{(0)} e)) = 0 \\
26. & \alpha^T (B^{(0)} (B^{(1)} e)) = 0 \\
27. & \beta^{(2)T} A^{(1)} e = 0 \\
28. & \beta^{(1)T} ((A^{(1)} e)(B^{(1)} e)) = 0 \\
29. & \beta^{(3)T} ((A^{(2)} e)(B^{(2)} e)) = 0 \\
30. & \beta^{(4)T} (A^{(2)} (B^{(0)} e)) = 0 \\
31. & \beta^{(2)T} (A^{(1)} (B^{(0)} e)) = 0 \\
32. & \beta^{(4)T} ((B^{(2)} e)^2 (A^{(2)} e)) = 0 \\
33. & \beta^{(4)T} (A^{(2)} (B^{(0)} e)^2) = 0 \\
34. & \beta^{(2)T} (A^{(1)} (B^{(0)} e)^2) = 0 \\
35. & \beta^{(1)T} (B^{(1)} (A^{(1)} e)) = 0 \\
36. & \beta^{(3)T} (B^{(2)} (A^{(1)} e)) = 0 \\
37. & \beta^{(2)T} (B^{(1)} e)^2 = 0 \\
38. & \beta^{(4)T} (B^{(2)} (B^{(1)} e)) = 0 \\
39. & \beta^{(2)T} (B^{(1)} (B^{(1)} e)) = 0 \\
40. & \beta^{(1)T} (B^{(1)} e)^3 = 0 \\
41. & \beta^{(3)T} (B^{(2)} e)^3 = 0 \\
42. & \beta^{(1)T} (B^{(1)} (B^{(1)} e)^2) = 0 \\
43. & \beta^{(3)T} (B^{(2)} (B^{(1)} e)^2) = 0 \\
44. & \beta^{(4)T} (B^{(2)} e)^4 = 0 \\
45. & \beta^{(4)T} (B^{(2)} (B^{(1)} e))^2 = 0
\end{array}$$

$$\begin{array}{l}
46. \quad \beta^{(4)T} ((B^{(2)} e)(B^{(2)} (B^{(1)} e))) = 0 \\
47. \quad \alpha^T ((B^{(0)} e)(B^{(0)} (B^{(1)} e))) = 0 \\
48. \quad \beta^{(1)T} ((A^{(1)} (B^{(0)} e))(B^{(1)} e)) = 0 \\
49. \quad \beta^{(3)T} ((A^{(2)} (B^{(0)} e))(B^{(2)} e)) = 0 \\
50. \quad \beta^{(1)T} (A^{(1)} (B^{(0)} (B^{(1)} e))) = 0 \\
51. \quad \beta^{(3)T} (A^{(2)} (B^{(0)} (B^{(1)} e))) = 0 \\
52. \quad \beta^{(4)T} ((B^{(2)} (A^{(1)} e))(B^{(2)} e)) = 0 \\
53. \quad \beta^{(1)T} (B^{(1)} (A^{(1)} (B^{(0)} e))) = 0 \\
54. \quad \beta^{(3)T} (B^{(2)} (A^{(1)} (B^{(0)} e))) = 0 \\
55. \quad \beta^{(1)T} ((B^{(1)} e)(B^{(1)} (B^{(1)} e))) = 0 \\
56. \quad \beta^{(3)T} ((B^{(2)} e)(B^{(2)} (B^{(1)} e))) = 0 \\
57. \quad \beta^{(1)T} (B^{(1)} (B^{(1)} (B^{(1)} e))) = 0 \\
58. \quad \beta^{(4)T} ((B^{(2)} e)(B^{(2)} (B^{(1)} e)^2)) = 0 \\
59. \quad \beta^{(4)T} ((B^{(2)} e)(B^{(2)} (B^{(1)} (B^{(1)} e)))) = 0
\end{array}$$

are fulfilled and if  $c^{(i)} = A^{(i)}e$  for  $i = 0, 1, 2$ , then the stochastic Runge–Kutta method (11) attains order 2.0 for the weak approximation of the solution of the Itô SDE (1).

**Remark 5.2** Due to Theorem 5.1 we have to solve 59 equations for  $m > 1$ . However in the case of  $m = 1$  the 59 conditions reduce to 28 conditions (see also [13]). For an explicit SRK method of type (11)  $s \geq 3$  is needed.

**Proof.** Firstly, we have to prove that the SRK method (11) is contained in the general class of SRK methods (3). Therefore, we choose  $\mathcal{M} = \{(k), (k, l) : 0 \leq k, l \leq m\}$  and define

$$\gamma_i^{(\iota)^{(k,\nu)}} \theta_\iota(h_n) = \begin{cases} \beta_i^{(1)} \hat{I}_{(k)} & \text{if } 0 < \iota = k, \nu = 0 \\ \beta_i^{(2)} \frac{\hat{I}_{(k,k)}}{\sqrt{h_n}} & \text{if } \iota = (k, k), 0 < k, \nu = 0 \\ \beta_i^{(3)} \hat{I}_{(k)} & \text{if } 0 < \iota = k, \nu = 1 \\ \beta_i^{(4)} \sqrt{h_n} & \text{if } 0 = \iota < k, \nu = 1 \\ 0 & \text{otherwise} \end{cases}$$

$$A_{ij}^{(k,\nu)(0,0)} h_n = \begin{cases} A_{ij}^{(0)} h_n & \text{if } k = \nu = 0 \\ A_{ij}^{(1)} h_n & \text{if } k > 0, \nu = 0 \\ A_{ij}^{(2)} h_n & \text{if } k > 0, \nu = 1 \\ 0 & \text{otherwise} \end{cases}$$

$$B_{ij}^{(\iota)^{(k,\nu)(r,\mu)}} \theta_\iota(h_n) = \begin{cases} B_{ij}^{(0)} \hat{I}_{(r)} & \text{if } k = \nu = \mu = 0, \iota = r \\ B_{ij}^{(1)} \sqrt{h_n} & \text{if } 0 = \iota < k = r, \nu = \mu = 0 \\ B_{ij}^{(2)} \frac{\hat{I}_{(k,r)}}{\sqrt{h_n}} & \text{if } \iota = (k, r), 0 < k \neq r, \nu = 1, \mu = 0 \\ 0 & \text{otherwise} \end{cases}$$

for  $k = 0, 1, \dots, m, r = 1, \dots, m$  and  $\iota, \nu, \mu \in \mathcal{M}$ . As a result of this we get

$$z_i^{(0,0)} = \alpha_i h_n \quad z_i^{(k,0)} = \beta_i^{(1)} \hat{I}_{(k)} + \beta_i^{(2)} \frac{\hat{I}_{(k,k)}}{\sqrt{h_n}} \quad z_i^{(k,1)} = \beta_i^{(3)} \hat{I}_{(k)} + \beta_i^{(4)} \sqrt{h_n}$$

$$\begin{aligned} Z_{ij}^{(0,0)(0,0)} &= A_{ij}^{(0)} h_n & Z_{ij}^{(k,0)(0,0)} &= A_{ij}^{(1)} h_n & Z_{ij}^{(k,1)(0,0)} &= A_{ij}^{(2)} h_n \\ Z_{ij}^{(0,0)(k,0)} &= B_{ij}^{(0)} \hat{I}_{(k)} & Z_{ij}^{(k,0)(k,0)} &= B_{ij}^{(1)} \sqrt{h_n} & Z_{ij}^{(k,1)(l,0)} &= B_{ij}^{(2)} \frac{\hat{I}_{(k,l)}}{\sqrt{h_n}} \end{aligned}$$

for  $1 \leq k, l \leq m$  with  $k \neq l$ . Further, we have  $H_i^{(0,0)} = H_i^{(0)}$ ,  $H_i^{(k,0)} = H_i^{(k)}$  and  $H_i^{(k,1)} = \hat{H}_i^{(k)}$ . Now, apply Theorem 4.2 for all  $\mathbf{t} \in TS(\Delta)$  with  $\rho(\mathbf{t}) \leq 2.5$ . We refer to [12] for all necessary trees and corresponding parameters  $\alpha_I(\mathbf{t})$ ,  $\alpha_\Delta(\mathbf{t})$  or  $\beta(\mathbf{t})$ . Apart from (13) and (14), the following moments are helpful in the subsequent calculations:  $E(\hat{I}_{(k,k)}^2) = \frac{1}{2}h^2$ ,  $E(\hat{I}_{(k)}^2 \hat{I}_{(k,k)}) = h^2$ ,  $E(\hat{I}_{(k,k)}^3) = \frac{1}{4}h^3$ ,  $E(\hat{I}_{(k)} \hat{I}_{(l)} \hat{I}_{(k,l)}) = \frac{1}{2}h^2$ ,  $E(\hat{I}_{(k,l)}^2) = \frac{1}{2}h^2$ ,  $E(\hat{I}_{(k,l)}^q) = 0$ ,  $E(\hat{I}_{(k)}^2 \hat{I}_{(k,l)}^2) = h^3$  and  $E(\hat{I}_{(k,l)} \hat{I}_{(l,k)}) = 0$  for  $q = 1, 3$  and  $k \neq l$ . In the following holds  $\beta(\mathbf{t}) = 1$  if not

stated otherwise.

Order 0.5 trees.

$$\mathbf{t}_{0.5.1} = (\sigma_{j_1}): \Phi_S(\mathbf{t}) = z^{(j_1,0)T} e + z^{(j_1,1)T} e$$

$$\text{With } \alpha_I(\mathbf{t}) = 0 \text{ follows } E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow \beta^{(4)T} e \sqrt{h} = 0.$$

In the following, we assume that condition 2. of Theorem 5.1 holds.

Order 1.0 trees.

$$\mathbf{t}_{1.1} = (\tau): \Phi_S(\mathbf{t}) = z^{(0,0)T} e$$

$$\text{With } \alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1 \text{ follows } E(\Phi_S(\mathbf{t})) = h \Leftrightarrow \alpha^T e h = h.$$

$$\mathbf{t}_{1.2} = (\sigma_{j_1}, \sigma_{j_2}): \Phi_S(\mathbf{t}) = (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} e + z^{(j_2,1)T} e)$$

$$\text{For } j_1 = j_2 \text{ with } \alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1 \text{ follows } E(\Phi_S(\mathbf{t})) = h \Leftrightarrow (\beta^{(1)T} e)^2 E(\hat{I}_{(j_1)}^2) + (\beta^{(2)T} e)^2 E(\hat{I}_{(j_1, j_1)}^2) h^{-1} + (\beta^{(3)T} e)^2 E(\hat{I}_{(j_1)}^2) + (\beta^{(1)T} e) (\beta^{(3)T} e) E(\hat{I}_{(j_1)}^2) = h.$$

$$\mathbf{t}_{1.3} = (\{\sigma_{j_2}\}_{j_1}): \Phi_S(\mathbf{t}) = z^{(j_1,0)T} Z^{(j_1,0)(j_2,0)} e + z^{(j_1,1)T} Z^{(j_1,1)(j_2,0)} e$$

$$\text{It holds } \alpha_I(\mathbf{t}) = 0 \text{ and } E(\Phi_S(\mathbf{t})) = 0.$$

Now, we additionally assume that condition 1. of Theorem 5.1 holds.

Order 1.5 trees.

$$\mathbf{t}_{1.5.2} = (\{\tau\}_{j_1}): \Phi_S(\mathbf{t}) = z^{(j_1,0)T} Z^{(j_1,0)(0,0)} e + z^{(j_1,1)T} Z^{(j_1,1)(0,0)} e$$

$$\text{With } \alpha_I(\mathbf{t}) = 0 \text{ follows } E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow \beta^{(4)T} A^{(2)} e h^{3/2} = 0.$$

$$\mathbf{t}_{1.5.4} = (\sigma_{j_1}, \sigma_{j_2}, \sigma_{j_3}):$$

$$\Phi_S(\mathbf{t}) = (z^{(j_1,0)T} e + z^{(j_1,1)T} e) (z^{(j_2,0)T} e + z^{(j_2,1)T} e) (z^{(j_3,0)T} e + z^{(j_3,1)T} e)$$

$$\text{For } j_1 = j_2 = j_3 \text{ with } \alpha_I(\mathbf{t}) = 0 \text{ follows } E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow$$

$$(\beta^{(1)T} e + \beta^{(3)T} e)^2 (\beta^{(2)T} e) E(\hat{I}_{(j_1)}^2 \hat{I}_{(j_1, j_1)}) h^{-1/2} + (\beta^{(2)T} e)^3 E(\hat{I}_{(j_1, j_1)}^3) h^{-3/2} = 0$$

$$\Leftrightarrow \beta^{(2)T} e = 0. \text{ In the following, due to } \mathbf{t}_{1.2} \text{ we assume } (\beta^{(1)T} e)^2 = 1 \text{ and } \beta^{(3)T} e = 0.$$

$$\mathbf{t}_{1.5.5} = (\{\sigma_{j_2}\}_{j_1}, \sigma_{j_3}):$$

$$\Phi_S(\mathbf{t}) = (z^{(j_1,0)T} Z^{(j_1,0)(j_2,0)} e + z^{(j_1,1)T} Z^{(j_1,1)(j_2,0)} e) (z^{(j_3,0)T} e + z^{(j_3,1)T} e)$$

$$\text{Case A): For } j_1 = j_2 = j_3 \text{ with } \alpha_I(\mathbf{t}) = 0 \text{ follows } E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow$$

$$(\beta^{(1)T} e) (\beta^{(1)T} B^{(1)} e) E(\hat{I}_{(j_1)}^2) \sqrt{h} = 0. \text{ Case B): For } j_1 \neq j_2 = j_3 \text{ with } \alpha_I(\mathbf{t}) = 0$$

$$\text{follows } E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow (\beta^{(1)T} e) (\beta^{(3)T} B^{(2)} e) E(\hat{I}_{(j_1)} \hat{I}_{(j_2)} \hat{I}_{(j_1, j_2)}) h^{-1/2} = 0.$$

$$\mathbf{t}_{1.5.6} = (\{\sigma_{j_2}, \sigma_{j_3}\}_{j_1}):$$

$$\begin{aligned} \Phi_S(\mathbf{t}) &= z^{(j_1,0)T} ((Z^{(j_1,0)(j_2,0)} e)(Z^{(j_1,0)(j_3,0)} e)) \\ &\quad + z^{(j_1,1)T} ((Z^{(j_1,1)(j_2,0)} e)(Z^{(j_1,1)(j_3,0)} e)) \end{aligned}$$

For  $j_1 \neq j_2 = j_3$  with  $\alpha_I(\mathbf{t}) = 0$  follows  $E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow \beta^{(4)T} (B^{(2)}e)^2 E(\hat{I}_{(j_1, j_2)}^2) h^{-1/2} = 0$ .

For the trees  $\mathbf{t}_{1.5.1} = ([\sigma_{j_1}])$ ,  $\mathbf{t}_{1.5.3} = (\tau, \sigma_{j_1})$  and  $\mathbf{t}_{1.5.7} = (\{\{\sigma_{j_3}\}_{j_2}\}_{j_1})$  holds  $\alpha_I(\mathbf{t}) = 0$  and  $E(\Phi_S(\mathbf{t})) = 0$ .

Order 2.0 trees:

$\mathbf{t}_{2.1} = ([\tau])$ :  $\Phi_S(\mathbf{t}) = z^{(0,0)T} Z^{(0,0)(0,0)} e$

With  $\alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow \alpha^T A^{(0)} e h^2 = \frac{1}{2}h^2$ .

$\mathbf{t}_{2.2} = (\tau, \tau)$ :  $\Phi_S(\mathbf{t}) = (z^{(0,0)T} e)^2$

With  $\alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1$  follows  $E(\Phi_S(\mathbf{t})) = h^2 \Leftrightarrow (\alpha^T e)^2 h^2 = h^2$ .

$\mathbf{t}_{2.4} = ([\sigma_{j_1}, \sigma_{j_2}])$ :  $\Phi_S(\mathbf{t}) = z^{(0,0)T} ((Z^{(0,0)(j_1,0)} e)(Z^{(0,0)(j_2,0)} e))$

For  $j_1 = j_2$  with  $\alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow \alpha^T (B^{(0)(j_1)} e)^2 E(\hat{I}_{(j_1)}^2) h = \frac{1}{2}h^2$ .

$\mathbf{t}_{2.5} = (\sigma_{j_1}, [\sigma_{j_2}])$ :  $\Phi_S(\mathbf{t}) = (z^{(j_1,0)T} e + z^{(j_1,1)T} e) (z^{(0,0)T} Z^{(0,0)(j_2,0)} e)$

For  $j_1 = j_2$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} e) (\alpha^T B^{(0)} e) E(\hat{I}_{(j_1)}^2) h = \frac{1}{2}h^2$ .

$\mathbf{t}_{2.7} = (\sigma_{j_1}, \sigma_{j_2}, \tau)$ :  $\Phi_S(\mathbf{t}) = (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} e + z^{(j_2,1)T} e)(z^{(0,0)T} e)$

For  $j_1 = j_2$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = h^2 \Leftrightarrow (\beta^{(1)T} e)^2 (\alpha^T e) E(\hat{I}_{(j_1)}^2) h = h^2$ .

$\mathbf{t}_{2.8} = (\sigma_{j_1}, \{\tau\}_{j_2})$ :

$\Phi_S(\mathbf{t}) = (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} Z^{(j_2,0)(0,0)} e + z^{(j_2,1)T} Z^{(j_2,1)(0,0)} e)$

For  $j_1 = j_2$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} e)(\beta^{(1)T} A^{(1)} e + \beta^{(3)T} A^{(2)} e) E(\hat{I}_{(j_1)}^2) h = \frac{1}{2}h^2$ .

Here, we claim that  $(\beta^{(1)T} e)(\beta^{(1)T} A^{(1)} e) = \frac{1}{2}$  and  $\beta^{(3)T} A^{(2)} e = 0$ .

$\mathbf{t}_{2.11} = (\sigma_{j_1}, \sigma_{j_2}, \sigma_{j_3}, \sigma_{j_4})$ :

$$\begin{aligned} \Phi_S(\mathbf{t}) &= (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} e + z^{(j_2,1)T} e) \\ &\quad \times (z^{(j_3,0)T} e + z^{(j_3,1)T} e)(z^{(j_4,0)T} e + z^{(j_4,1)T} e) \end{aligned}$$

Case A): For  $j_1 = j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = \beta(\mathbf{t}) = 1$  follows  $E(\Phi_S(\mathbf{t})) = 3h^2 \Leftrightarrow (\beta^{(1)T} e)^4 E(\hat{I}_{(j_1)}^4) = 3h^2$ . Case B): For  $j_1 = j_2 \neq j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = \alpha_\Delta(\mathbf{t}) = 1$  and  $\beta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = h^2 \Leftrightarrow (\beta^{(1)T} e)^2 (\beta^{(1)T} e)^2 E(\hat{I}_{(j_1)}^2 \hat{I}_{(j_3)}^2) = h^2$ .

$$\mathbf{t}_{2.12} = (\sigma_{j_1}, \sigma_{j_2}, \{\sigma_{j_4}\}_{j_3}):$$

$$\begin{aligned} \Phi_S(\mathbf{t}) &= (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} e + z^{(j_2,1)T} e) \\ &\quad \times (z^{(j_3,0)T} Z^{(j_3,0)(j_4,0)} e + z^{(j_3,1)T} Z^{(j_3,1)(j_4,0)} e) \end{aligned}$$

Case A): For  $j_1 = j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 4$ ,  $\alpha_\Delta(\mathbf{t}) = 6$  and  $\beta(\mathbf{t}) = 1$  follows  $E(\Phi_S(\mathbf{t})) = h^2 \Leftrightarrow (\beta^{(1)T} e)^2 (\beta^{(2)T} B^{(1)} e) E(\hat{I}_{(j_1)}^2 \hat{I}_{(j_1, j_1)}) = h^2$ . Case B): For  $j_1 = j_3 \neq j_2 = j_4$  with  $\alpha_I(\mathbf{t}) = 4$ ,  $\alpha_\Delta(\mathbf{t}) = 6$  and  $\beta(\mathbf{t}) = 2$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} e)^2 (\beta^{(4)T} B^{(2)} e) E(\hat{I}_{(j_1)} \hat{I}_{(j_2)} \hat{I}_{(j_1, j_2)}) = \frac{1}{2}h^2$ .

$$\mathbf{t}_{2.13} = (\sigma_{j_1}, \{\sigma_{j_3}, \sigma_{j_4}\}_{j_2}):$$

$$\begin{aligned} \Phi_S(\mathbf{t}) &= (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} ((Z^{(j_2,0)(j_3,0)} e)(Z^{(j_2,0)(j_4,0)} e)) \\ &\quad + z^{(j_2,1)T} ((Z^{(j_2,1)(j_3,0)} e)(Z^{(j_2,1)(j_4,0)} e))) \end{aligned}$$

Case A): For  $j_1 = j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 4$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} e)(\beta^{(1)T} (B^{(1)} e)^2) E(\hat{I}_{(j_1)}^2) h = \frac{1}{2}h^2$ . Case B): For  $j_1 = j_2 \neq j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 4$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} e)(\beta^{(3)T} (B^{(2)} e)^2) E(\hat{I}_{(j_1)}^2 \hat{I}_{(j_1, j_3)}^2) h^{-1} = \frac{1}{2}h^2$ .

$$\mathbf{t}_{2.14} = (\sigma_{j_1}, \{\{\sigma_{j_4}\}_{j_3}\}_{j_2}):$$

$$\begin{aligned} \Phi_S(\mathbf{t}_{2.14}) &= (z^{(j_1,0)T} e + z^{(j_1,1)T} e)(z^{(j_2,0)T} (Z^{(j_2,0)(j_3,0)} (Z^{(j_3,0)(j_4,0)} e)) \\ &\quad + z^{(j_2,1)T} (Z^{(j_2,1)(j_3,0)} (Z^{(j_3,0)(j_4,0)} e))) \end{aligned}$$

Case A): For  $j_1 = j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 0$  follows  $E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow (\beta^{(1)T} e)(\beta^{(1)T} (B^{(1)} (B^{(1)} e))) E(\hat{I}_{(j_1)}^2) h = 0$ . Case B): For  $j_1 = j_3 = j_4 \neq j_2$  with  $\alpha_I(\mathbf{t}) = 0$  follows  $E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow (\beta^{(1)T} e)(\beta^{(3)T} (B^{(2)} (B^{(1)} e))) E(\hat{I}_{(j_1)} \hat{I}_{(j_2)} \hat{I}_{(j_2, j_1)}) = 0$ .

$$\mathbf{t}_{2.15} = (\{\sigma_{j_2}\}_{j_1}, \{\sigma_{j_4}\}_{j_3}):$$

$$\begin{aligned} \Phi_S(\mathbf{t}) &= (z^{(j_1,0)T} Z^{(j_1,0)(j_2,0)} e + z^{(j_1,1)T} Z^{(j_1,1)(j_2,0)} e)(z^{(j_3,0)T} Z^{(j_3,0)(j_4,0)} e \\ &\quad + z^{(j_3,1)T} Z^{(j_3,1)(j_4,0)} e) \end{aligned}$$

Case A): For  $j_1 = j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(1)T} B^{(1)} e)^2 E(\hat{I}_{(j_1)}^2) h + (\beta^{(2)T} B^{(1)} e)^2 E(\hat{I}_{(j_1, j_1)}^2) = \frac{1}{2}h^2$ . Case B): For  $j_1 = j_3 \neq j_2 = j_4$  with  $\alpha_I(\mathbf{t}) = 2$  and  $\alpha_\Delta(\mathbf{t}) = 3$  follows  $E(\Phi_S(\mathbf{t})) = \frac{1}{2}h^2 \Leftrightarrow (\beta^{(4)T} B^{(2)} e)^2 E(\hat{I}_{(j_1, j_2)}^2) + (\beta^{(3)T} B^{(2)} e)^2 E(\hat{I}_{(j_1)}^2 \hat{I}_{(j_1, j_2)}^2) h^{-1} = \frac{1}{2}h^2$ .

$\mathbf{t}_{2.17} = (\{\sigma_{j_2}, \{\sigma_{j_4}\}_{j_3}\}_{j_1})$ :

$$\begin{aligned}\Phi_S(\mathbf{t}) &= z^{(j_1,0)T}((Z^{(j_1,0)(j_2,0)}e)(Z^{(j_1,0)(j_3,0)}(Z^{(j_3,0)(j_4,0)}e))) \\ &\quad + z^{(j_1,1)T}((Z^{(j_1,1)(j_2,0)}e)(Z^{(j_1,1)(j_3,0)}(Z^{(j_3,0)(j_4,0)}e)))\end{aligned}$$

For  $j_1 \neq j_2 = j_3 = j_4$  with  $\alpha_I(\mathbf{t}) = 0$  follows  $E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow \beta^{(4)T}((B^{(2)}e)(B^{(2)}(B^{(1)}e))) E(\hat{I}_{(j_1, j_2)}^2) = 0$ .

$\mathbf{t}_{2.20} = (\{[\sigma_{j_2}]\}_{j_1})$ :

$$\Phi_S(\mathbf{t}) = z^{(j_1,0)T}(Z^{(j_1,0)(0,0)}(Z^{(0,0)(j_2,0)}e)) + z^{(j_1,1)T}(Z^{(j_1,1)(0,0)}(Z^{(0,0)(j_2,0)}e))$$

For  $j_1 = j_2$  with  $\alpha_I(\mathbf{t}) = 0$  follows  $E(\Phi_S(\mathbf{t})) = 0 \Leftrightarrow \beta^{(3)T}(A^{(2)}(B^{(0)}e)) E(\hat{I}_{(j_1)}^2) + \beta^{(1)T}(A^{(1)}(B^{(0)}e)) E(\hat{I}_{(j_1)}^2) = 0$ . Here, we claim that  $\beta^{(3)T}(A^{(2)}(B^{(0)}e)) = 0$  and  $\beta^{(1)T}(A^{(1)}(B^{(0)}e)) = 0$ .

For all correlations between  $j_1, \dots, j_4$  which have not been considered explicitly holds  $\alpha_I(\mathbf{t}) = 0$  and  $E(\Phi_S(\mathbf{t})) = 0$ . Further, for the trees  $\mathbf{t}_{2.3} = (\{[\sigma_{j_2}]\}_{j_1})$ ,  $\mathbf{t}_{2.6} = (\{\sigma_{j_2}\}_{j_1}, \tau)$ ,  $\mathbf{t}_{2.9} = (\{\{\tau\}_{j_2}\}_{j_1})$ ,  $\mathbf{t}_{2.10} = (\{\sigma_{j_2}, \tau\}_{j_1})$ ,  $\mathbf{t}_{2.16} = (\{\sigma_{j_2}, \sigma_{j_3}, \sigma_{j_4}\}_{j_1})$ ,  $\mathbf{t}_{2.18} = (\{\{\sigma_{j_3}, \sigma_{j_4}\}_{j_2}\}_{j_1})$  and  $\mathbf{t}_{2.19} = (\{\{\{\sigma_{j_4}\}_{j_3}\}_{j_2}\}_{j_1})$  holds  $\alpha_I(\mathbf{t}) = 0$  and  $E(\Phi_S(\mathbf{t})) = 0$ .

Finally, we have to consider all trees  $\mathbf{t} \in TS(\Delta)$  with  $\rho(\mathbf{t}) = 2.5$  for which due to  $\alpha_I(\mathbf{t}) = 0$  the condition  $E(\Phi_S(\mathbf{t})) = 0$  has to be fulfilled. Since the calculations are analogous to the ones already performed, repetition is avoided (see [12] for all trees up to order 2.5). Leaving out the trees which don't supply any new restrictions, we calculate the following conditions:

Table 1: Conditions from  $\mathbf{t} \in TS(\Delta)$  with  $\rho(\mathbf{t}) = 2.5$ .

$\mathbf{t}$	correlation	condition
$(\{\tau, \tau\}_{j_1})$		24.
$(\{[\tau]\}_{j_1})$		25.
$(\{[\sigma_{j_2}]\}_{j_1}, \sigma_{j_3})$	$j_1 = j_2 = j_3$	26.
$([\sigma_{j_1}, \{\sigma_{j_3}\}_{j_2}])$	$j_1 = j_2 = j_3$	47.
$(\{\tau\}_{j_1}, \sigma_{j_2}, \sigma_{j_3})$	$j_1 = j_2 = j_3$	27. + 7.
	$j_1 \neq j_2 = j_3$	7.
$(\{\tau, \sigma_{j_2}\}_{j_1}, \sigma_{j_3})$	$j_1 = j_2 = j_3$	28.
	$j_1 \neq j_2 = j_3$	29.
$(\{[\sigma_{j_2}]\}_{j_1}, \sigma_{j_3})$	$j_1 = j_2 = j_3$	30. + 31.

$\mathbf{t}$	correlation	condition
	$j_1 \neq j_2 = j_3$	30.
$(\{\tau, \sigma_{j_2}, \sigma_{j_3}\}_{j_1})$	$j_1 \neq j_2 = j_3$	32.
$(\{[\sigma_{j_2}], \sigma_{j_3}\}_{j_1})$	$j_1 = j_2 = j_3$	48.
	$j_1 \neq j_2 = j_3$	49.
$(\{[\sigma_{j_2}, \sigma_{j_3}]\}_{j_1})$	$j_1 = j_2 = j_3$	33. + 34.
	$j_1 \neq j_2 = j_3$	33.
$(\{\{[\sigma_{j_3}]_{j_2}\}\}_{j_1})$	$j_1 = j_2 = j_3$	50. + 51.
$(\{\{\tau\}_{j_2}\}_{j_1}, \sigma_{j_3})$	$j_1 = j_2 = j_3$	35.
	$j_1 \neq j_2 = j_3$	36.
$(\{\{\tau\}_{j_2}, \sigma_{j_3}\}_{j_1})$	$j_1 \neq j_2 = j_3$	52.
$(\{\{[\sigma_{j_3}]\}_{j_2}\}_{j_1})$	$j_1 = j_2 = j_3$	53.
	$j_1 \neq j_2 = j_3$	54.
$(\sigma_{j_1}, \sigma_{j_2}, \{\sigma_{j_4}, \sigma_{j_5}\}_{j_3})$	$j_1 = j_2 = j_3 = j_4 = j_5$	37.
	$j_1 = j_4 \neq j_3, j_2 = j_5 \neq j_3$	9.
$(\sigma_{j_1}, \sigma_{j_2}, \{\{\sigma_{j_5}\}_{j_4}\}_{j_3})$	$j_1 = j_2 = j_3 = j_4 = j_5$	39.
	$j_1 = j_3 \neq j_2 = j_4 = j_5$	38.
$(\{\sigma_{j_2}, \sigma_{j_3}, \sigma_{j_4}\}_{j_1}, \sigma_{j_5})$	$j_1 = j_2 = j_3 = j_4 = j_5$	40.
	$j_1 \neq j_2 = j_3, j_1 \neq j_4 = j_5, j_3 \neq j_4$	41.
	$j_1 \neq j_2 = j_3 = j_4 = j_5$	41.
$(\{\sigma_{j_2}, \{\sigma_{j_4}\}_{j_3}\}_{j_1}, \sigma_{j_5})$	$j_1 = j_2 = j_3 = j_4 = j_5$	55.
	$j_1 = j_5 \neq j_2 = j_3 = j_4$	56.
$(\{\{\sigma_{j_3}, \sigma_{j_4}\}_{j_2}\}_{j_1}, \sigma_{j_5})$	$j_1 = j_2 = j_3 = j_4 = j_5$	42.
	$j_1 \neq j_2 = j_3 = j_4 = j_5$	43.
$(\{\{\{\sigma_{j_4}\}_{j_3}\}_{j_2}\}_{j_1}, \sigma_{j_5})$	$j_1 = j_2 = j_3 = j_4 = j_5$	57.
	$j_1 \neq j_2 = j_3 = j_4 = j_5$	21.
$(\{\sigma_{j_2}, \sigma_{j_3}, \sigma_{j_4}, \sigma_{j_5}\}_{j_1})$	$j_1 \neq j_2 = j_3, j_1 \neq j_4 = j_5$	44.
$(\{\sigma_{j_2}, \{\sigma_{j_4}, \sigma_{j_5}\}_{j_3}\}_{j_1})$	$j_1 \neq j_2 = j_3 = j_4 = j_5$	58.
$(\{\sigma_{j_2}, \{\{\sigma_{j_5}\}_{j_4}\}_{j_3}\}_{j_1})$	$j_1 \neq j_2 = j_3 = j_4 = j_5$	59.



$\mathbf{t}$	correlation	condition
$(\{\{\sigma_{j_3}\}_{j_2}, \{\sigma_{j_5}\}_{j_4}\}_{j_1})$	$j_1 \neq j_2 = j_3 = j_4 = j_5$	45.

Now, we just have to summarize the calculated conditions in order to arrive at the conditions in Theorem 5.1. Finally, the approximation  $Y$  by the SRK method (11) has uniformly bounded moments due to  $E(z^{(k,\nu)T}e) = 0$  for  $1 \leq k \leq m$  and  $\nu \in \{0, 1\}$ .  $\square$

Considering the order conditions 1.–9. of Theorem 5.1, we can easily calculate SRK schemes converging with order 1.0 in the weak sense. For example, the well known Euler-Maruyama scheme (see, e.g., [5]) belongs to the introduced class of SRK methods having order 1.0 with  $s = 1$  stage and with coefficients  $\alpha_1 = \beta_1^{(1)} = 1$ ,  $\beta_1^{(2)} = \beta_1^{(3)} = \beta_1^{(4)} = 0$ ,  $A_{11}^{(0)} = A_{11}^{(1)} = 0$  and  $B_{11}^{(0)} = B_{11}^{(1)} = 0$ . Further, if we calculate order 2.0 SRK schemes with  $s \geq 3$  stages then there are some degrees of freedom in choosing the coefficients. Especially, it is possible to calculate a SRK scheme converging with some higher order if it is applied to a deterministic ordinary differential equation. For example, if the weights  $\alpha_i$  and the coefficients  $A_{ij}^{(0)}$  are chosen such that the conditions of Theorem 5.1 and additionally the conditions  $\alpha^T(A^{(0)}(A^{(0)}e)) = \frac{1}{6}$  and  $\alpha^T(A^{(0)}e)^2 = \frac{1}{3}$  are fulfilled (see, e.g. [3]), then the SRK scheme is of order three in the case of  $b \equiv 0$  in SDE (1). Therefore, let  $(p_D, p_S)$  with  $p_D \geq p_S$  denote the order of convergence of the SRK scheme if it is applied to a deterministic or stochastic differential equation, respectively. Thus, the scheme converges at least with order  $p = p_S$  in the weak sense and we suppose better convergence for schemes with  $p_D > p_S$ , particularly for SDEs with small noise. The SRK schemes

0									
$\frac{2}{3}$	$\frac{2}{3}$		1						
$\frac{2}{3}$	$-\frac{1}{3}$	1	0	0					
0									
1	1		1						
1	1	0	-1	0					
0									
0	0		1						
0	0	0	-1	0					
	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$
				$-\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$
0									
1	1		$\frac{3-2\sqrt{6}}{5}$						
$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	$\frac{6+\sqrt{6}}{10}$	0					
0									
1	1		1						
1	1	0	-1	0					
0									
0	0		1						
0	0	0	-1	0					
	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{2}{3}$	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$
				$-\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$

Table 2  
SRK scheme RI1 and RI3 of order  $p_D = 3.0$  and  $p_S = 2.0$ .

0									
1	1		$\frac{1}{3}$						
$\frac{5}{12}$	$\frac{25}{144}$	$\frac{35}{144}$	$-\frac{5}{6}$	0					
0									
$\frac{1}{4}$	$\frac{1}{4}$		$\frac{1}{2}$						
$\frac{1}{4}$	$\frac{1}{4}$	0	$-\frac{1}{2}$	0					
0									
0	0		1						
0	0	0	-1	0					
	$\frac{1}{10}$	$\frac{3}{14}$	$\frac{24}{35}$	1	-1	-1	0	1	-1
			$\frac{1}{2}$	$-\frac{1}{4}$	$-\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$	

0									
1	1		1						
0	0	0	0	0					
0									
1	1		1						
1	1	0	-1	0					
0									
0	0		1						
0	0	0	-1	0					
	$\frac{1}{2}$	$\frac{1}{2}$	0	$\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$
			$-\frac{1}{2}$	$\frac{1}{4}$	$\frac{1}{4}$	0	$\frac{1}{2}$	$-\frac{1}{2}$	

Table 3

SRK scheme RI5 of order  $p_D = 3.0$  and  $p_S = 2.0$  and RI6 of order  $p_D = p_S = 2.0$ .

RI1, RI3 and RI5 presented in Table 2 and Table 3, respectively, are of order  $p_S = 2.0$  and  $p_D = 3.0$ , and the SRK scheme RI6 in Table 3 is of order  $p_D = p_S = 2.0$ . Remark, that the schemes RI1, RI3, RI5 and RI6 coincide with the SRK schemes RI1W1, RI3W1, RI5W1 and PL1W1, respectively, presented in [13] in the case that they are applied to SDEs with scalar noise, i.e. in the case of  $d \geq 1$  and  $m = 1$ , due to the choice of  $A_{ij}^{(2)} = 0$ .

## 6 Numerical Example

In order to verify the theoretical results, we compare the SRK scheme (RI6) with the order one Euler-Maruyama scheme (EM), with the second order SRK scheme (PL1WM) due to Platen [5] which is also contained in the class of SRK methods proposed in [19] and with the extrapolated Euler-Maruyama scheme (ExEu) due to Talay and Tubaro [18] attaining order two. The extrapolated Euler-Maruyama approximation is given by  $2 E(f(Z_T^{h/2})) - E(f(Z_T^h))$  based on the Euler-Maruyama approximations  $Z_T^{h/2}$  and  $Z_T^h$  calculated with step sizes  $h$  and  $h/2$ . In the following, we approximate the values  $E(f(X_T))$  with  $f(x^1, x^2) = x^1$  and  $f(x^1, x^2) = x^1 x^2$  or  $f(x^1, x^2) = (x^1)^2$  by Monte Carlo simulation. Therefore, we estimate  $E(f(Y_T))$  by the sample average of  $M$  independent simulated realizations of the approximations  $f(Y_{T,k})$ ,  $k = 1, \dots, M$ , with  $Y_{T,k}$  calculated by the scheme under consideration. Then, the error is denoted by  $\hat{\mu} = E(f(X_T)) - \frac{1}{M} \sum_{k=1}^M f(Y_{T,k})$ . The empirical variance  $\hat{\sigma}_\mu^2$  of the error  $\hat{\mu}$  is calculated following [5] based on  $M_1$  batches with  $M_2$  trajectories in each, i.e.,  $M = M_1 \cdot M_2$ . Since we want to investigate the systematic error of the considered schemes, we have to minimize the statistical error by choosing  $M$  very large [5], i.e. much larger than usually necessary for an approxima-

tion in practice. The obtained errors at time  $T = 1.0$  are plotted versus the corresponding step sizes with double logarithmic scale in order to analyze the empirical order of convergence. In the presented figures, reference lines with slope 1.0 and 2.0 are plotted for comparison.

As the first test equation, we consider for  $d = m = 2$  the linear SDE system with commutative noise

$$d \begin{pmatrix} X_t^1 \\ X_t^2 \end{pmatrix} = \begin{pmatrix} \frac{3}{2}X_t^1 \\ \frac{3}{2}X_t^2 \end{pmatrix} dt + \begin{pmatrix} \frac{1}{10}X_t^1 \\ 0 \end{pmatrix} dW_t^1 + \begin{pmatrix} 0 \\ \frac{1}{10}X_t^2 \end{pmatrix} dW_t^2, \quad (15)$$

with initial value  $X_0 = (\frac{1}{10}, \frac{1}{10})^T$ . Then, the moments of the solution are given as  $E(X_t^i) = \frac{1}{10} \exp(\frac{3}{2}t)$  and  $E((X_t^i)^2) = \frac{1}{100} \exp(\frac{301}{100}t)$  for  $i = 1, 2$ . Here, we choose  $M_1 = 20$  batches with  $M_2 = 5 \times 10^6$  trajectories in each and consider the step sizes  $2^0, \dots, 2^{-7}$ . The errors  $|\hat{\mu}|$  and empirical variances  $\hat{\sigma}_\mu^2$  with corresponding step sizes are presented in Fig. 2 and Tab. 4–5.

Fig. 2. Step size vs. error for the approximation of  $E(X_T^1)$  in the left and of  $E((X_T^1)^2)$  in the right figure for SDE (15).

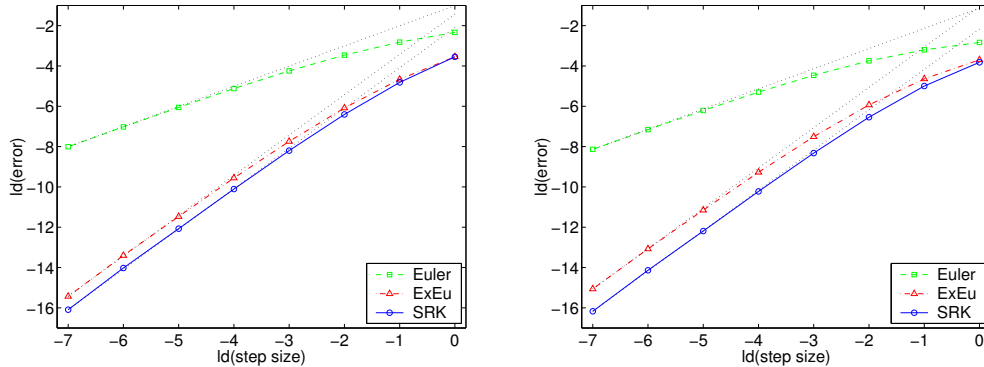


Table 4

Results for the approximation of  $E(X_t^1)$  for SDE (15).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	8.57e-2	1.21e-10	1.98e-1	1.95e-11	8.57e-2	1.19e-10	8.57e-02	1.21e-10
$2^{-1}$	3.56e-2	1.85e-10	1.42e-1	4.40e-11	3.95e-2	1.72e-10	3.56e-02	1.85e-10
$2^{-2}$	1.18e-2	3.43e-10	9.07e-2	1.33e-10	1.48e-2	3.27e-10	1.18e-02	3.43e-10
$2^{-3}$	3.40e-3	1.71e-10	5.27e-2	9.67e-11	4.67e-3	1.64e-10	3.40e-03	1.71e-10
$2^{-4}$	9.08e-4	4.41e-10	2.87e-2	3.25e-10	1.33e-3	4.35e-10	9.08e-04	4.41e-10
$2^{-5}$	2.32e-4	8.51e-10	1.50e-2	7.25e-10	3.52e-4	8.43e-10	2.32e-04	8.51e-10
$2^{-6}$	6.01e-5	6.18e-10	7.69e-3	5.71e-10	9.23e-5	6.16e-10	6.01e-05	6.18e-10
$2^{-7}$	1.43e-5	4.18e-10	3.89e-3	4.02e-10	2.27e-5	4.17e-10	1.43e-05	4.18e-10

As a second test equation, we consider for  $d = m = 2$  the linear SDE system

Table 5

Results for the approximation of  $E((X_t^1)^2)$  for SDE (15).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	7.08e-2	6.32e-11	1.40e-1	4.84e-12	7.73e-2	5.13e-11	7.08e-02	6.32e-11
$2^{-1}$	3.14e-2	1.31e-10	1.09e-1	1.69e-11	4.01e-2	1.05e-10	3.14e-02	1.31e-10
$2^{-2}$	1.07e-2	2.72e-10	7.44e-2	6.97e-11	1.64e-2	2.40e-10	1.07e-02	2.72e-10
$2^{-3}$	3.12e-3	1.48e-10	4.54e-2	6.50e-11	5.50e-3	1.38e-10	3.12e-03	1.48e-10
$2^{-4}$	8.37e-4	3.74e-10	2.54e-2	2.40e-10	1.62e-3	3.65e-10	8.37e-04	3.74e-10
$2^{-5}$	2.14e-4	7.26e-10	1.35e-2	5.76e-10	4.39e-4	7.17e-10	2.14e-04	7.26e-10
$2^{-6}$	5.55e-5	5.06e-10	6.99e-3	4.51e-10	1.16e-4	5.04e-10	5.55e-05	5.06e-10
$2^{-7}$	1.36e-5	3.40e-10	3.55e-3	3.21e-10	2.92e-5	3.39e-10	1.36e-05	3.40e-10

with non-commutative noise

$$d \begin{pmatrix} X_t^1 \\ X_t^2 \end{pmatrix} = \begin{pmatrix} -\frac{1}{2}X_t^1 + X_t^2 \\ \frac{1}{2}X_t^1 \end{pmatrix} dt + \begin{pmatrix} \frac{\sqrt{3}}{2}X_t^1 - \frac{\sqrt{3}}{2}X_t^2 \\ 0 \end{pmatrix} dW_t^1 + \begin{pmatrix} \frac{1}{2}X_t^1 + \frac{1}{2}X_t^2 \\ X_t^1 \end{pmatrix} dW_t^2, \quad (16)$$

with initial value  $X_0 = (\frac{1}{10}, \frac{1}{10})^T$ . Then, we can calculate the first moment of  $X_t^i$  as  $E(X_t^i) = \frac{1}{10}e^{\frac{1}{2}t}$  and the second moment as well as the mixed second moment of  $X_t^1$  and  $X_t^2$  as  $E((X_t^i)^2) = E(X_t^1 X_t^2) = \frac{1}{100}e^{2t}$  for  $i = 1, 2$ . For the approximation, we choose  $M_1 = 20$  and  $M_2 = 5 \times 10^7$ . The results for step sizes  $2^0, \dots, 2^{-5}$  are presented in Figure 3 and Tables 6–7.

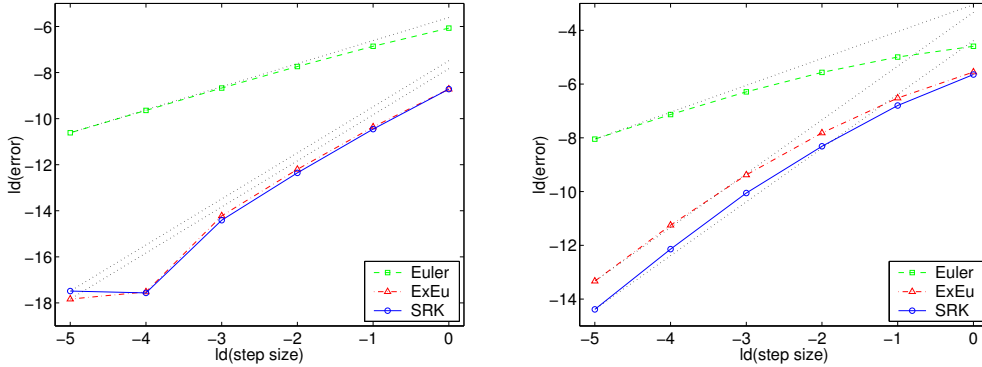
Fig. 3. Step size vs. error for the approximation of  $E(X_T^1)$  in the left and of  $E(X_T^1 X_T^2)$  in the right figure for SDE (16).

Table 6

Results for the approximation of  $E(X_t^1)$  for SDE (16).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	2.37e-3	8.12e-10	1.49e-2	3.82e-10	2.37e-3	8.90e-10	2.37e-03	8.12e-10
$2^{-1}$	7.12e-4	5.48e-10	8.62e-3	4.57e-10	7.60e-4	4.19e-10	7.12e-04	5.48e-10
$2^{-2}$	1.91e-4	6.40e-10	4.69e-3	4.97e-10	2.12e-4	7.57e-10	1.91e-04	6.40e-10
$2^{-3}$	4.59e-5	1.08e-09	2.45e-3	8.89e-10	5.24e-5	1.10e-09	4.59e-05	1.08e-09
$2^{-4}$	5.15e-6	8.81e-10	1.25e-3	7.57e-10	5.32e-6	9.40e-10	5.15e-06	8.81e-10
$2^{-5}$	5.45e-6	1.24e-09	6.38e-4	1.12e-09	4.28e-6	1.38e-09	5.45e-06	1.24e-09

The third test equation is a non-linear SDE system for  $d = m = 2$  with

Table 7

Results for the approximation of  $E(X_t^1 X_t^2)$  for SDE (16).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	2.00e-2	2.51e-10	4.14e-2	3.25e-11	2.13e-2	2.35e-10	2.00e-02	2.51e-10
$2^{-1}$	8.98e-3	7.43e-10	3.14e-2	5.53e-11	1.09e-2	3.84e-10	8.98e-03	7.43e-10
$2^{-2}$	3.14e-3	1.89e-09	2.11e-2	1.75e-10	4.43e-3	1.28e-09	3.14e-03	1.89e-09
$2^{-3}$	9.41e-4	4.41e-09	1.28e-2	1.24e-09	1.50e-3	4.01e-09	9.41e-04	4.41e-09
$2^{-4}$	2.21e-4	3.70e-09	7.12e-3	1.46e-09	4.10e-4	3.90e-09	2.21e-04	3.70e-09
$2^{-5}$	4.68e-5	7.70e-09	3.78e-3	4.33e-09	9.69e-5	7.61e-09	4.68e-05	7.70e-09

non-commutative noise given by

$$d \begin{pmatrix} X_t^1 \\ X_t^2 \end{pmatrix} = \begin{pmatrix} -\frac{1}{2}X_t^1 + \frac{3}{2}X_t^2 \\ \frac{3}{2}X_t^1 - \frac{1}{2}X_t^2 \end{pmatrix} dt + \begin{pmatrix} \sqrt{\frac{3}{4}(X_t^1)^2 - \frac{3}{2}X_t^1 X_t^2 + \frac{3}{4}(X_t^2)^2 + \frac{3}{20}} \\ 0 \end{pmatrix} dW_t^1 \\ + \begin{pmatrix} -\sqrt{\frac{1}{4}(X_t^1)^2 - \frac{1}{2}X_t^1 X_t^2 + \frac{1}{4}(X_t^2)^2 + \frac{1}{20}} \\ \sqrt{(X_t^1)^2 - 2X_t^1 X_t^2 + X_t^2 + \frac{1}{5}} \end{pmatrix} dW_t^2, \quad (17)$$

with initial value  $X_0 = (\frac{1}{10}, \frac{1}{10})^T$ . Then, the corresponding moments of the solution are  $E(X_t^i) = \frac{1}{10} \exp(t)$ ,  $E((X_t^i)^2) = \frac{3}{50} \exp(2t) - \frac{1}{10} \exp(-t) + \frac{1}{20}$  and  $E(X_t^1 X_t^2) = \frac{3}{50} \exp(2t) + \frac{1}{5} \exp(-t) - \frac{1}{4}$  for  $i = 1, 2$ . Here, we choose  $M_1 = 20$  and  $M_2 = 5 \times 10^7$ . The results are presented in Figure 4 and Tables 8–9.

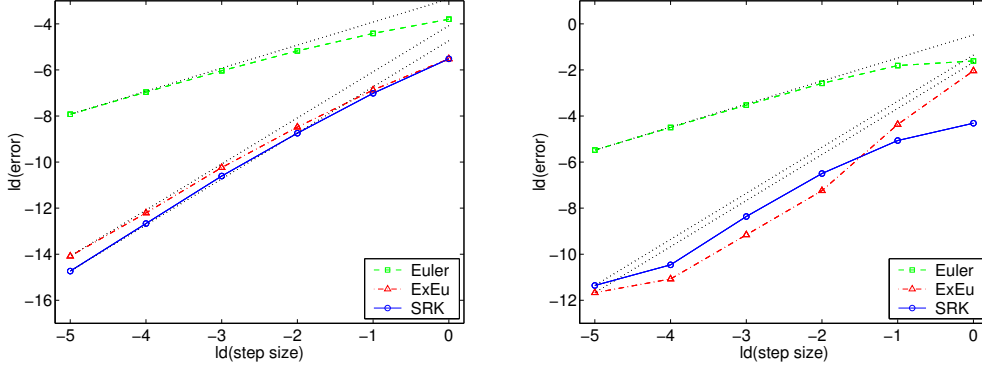
Fig. 4. Step size vs. error for the approximation of  $E(X_T^1)$  in the left and of  $E(X_T^1 X_T^2)$  in the right figure for SDE (17).

Table 8

Results for the approximation of  $E(X_t^1)$  for SDE (17).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	2.18e-2	3.88e-09	7.18e-2	2.98e-09	2.18e-2	1.31e-08	2.18e-02	4.57e-09
$2^{-1}$	7.74e-3	9.00e-09	4.68e-2	6.80e-09	8.53e-3	1.33e-08	7.76e-03	1.05e-08
$2^{-2}$	2.33e-3	1.41e-08	2.77e-2	1.07e-08	2.80e-3	1.64e-08	2.33e-03	1.50e-08
$2^{-3}$	6.41e-4	5.40e-09	1.53e-2	5.71e-09	8.28e-4	6.13e-09	6.38e-04	7.05e-09
$2^{-4}$	1.54e-4	6.57e-09	8.01e-3	6.95e-09	2.11e-4	7.79e-09	1.49e-04	7.53e-09
$2^{-5}$	3.67e-5	3.73e-09	4.12e-3	4.79e-09	5.77e-5	5.29e-09	3.16e-05	4.50e-09

For the SDEs (15) and (16), the results of SRK method RI6 coincide with

Table 9

Results for the approximation of  $E(X_t^1 X_t^2)$  for SDE (17).

$h$	RI6		EM		ExEu		PL1WM	
	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$	$ \hat{\mu} $	$\hat{\sigma}_\mu^2$
$2^{-0}$	5.03e-2	4.34e-09	3.27e-1	8.15e-10	2.43e-1	6.46e-09	3.92e-02	2.83e-09
$2^{-1}$	2.98e-2	6.24e-09	2.85e-1	4.33e-09	4.85e-2	1.45e-08	2.73e-02	5.34e-09
$2^{-2}$	1.11e-2	1.76e-08	1.67e-1	6.23e-09	6.60e-3	2.84e-08	1.12e-02	9.88e-09
$2^{-3}$	3.03e-3	7.73e-08	8.67e-2	2.19e-08	1.75e-3	1.05e-07	3.36e-03	8.05e-08
$2^{-4}$	7.09e-4	1.37e-07	4.42e-2	3.46e-08	4.63e-4	1.85e-07	8.18e-04	1.86e-07
$2^{-5}$	3.81e-4	3.16e-07	2.24e-2	1.31e-07	3.07e-4	3.08e-07	4.05e-04	3.07e-07

the results of the SRK scheme PL1WM since both schemes coincide for linear drift and diffusion functions. However, this is not the case for nonlinear SDEs like (17). It turns out that the SRK method RI6 yields often better results than the extrapolated Euler-Maruyama scheme while the empirical variances are assimilable.

## 7 Conclusion

The main advantage of the introduced class of SRK methods (11) is the significant reduction of the computational costs. Considering for example the order two SRK schemes proposed in [5] p.486 and in [19] for Itô SDEs with a multi-dimensional Wiener process, then at least 2 evaluations of the drift function  $a$  and  $2m + 1$  evaluations of each diffusion function  $b^j$ ,  $j = 1, \dots, m$ , are necessary for each step. Further,  $m(m + 1)/2$  independent random variables have to be simulated for the schemes in [5,19] for each step. Nearly the same holds for the class of SRK schemes recently proposed in [7] for Stratonovich SDEs which need 4 evaluations of the drift  $a$  and  $3m + 1$  evaluations of each diffusion function  $b^j$ ,  $j = 1, \dots, m$ , in each step. Due to the dependence of the computational costs on the dimension  $m$  of the driving Wiener process, these SRK methods are not of much relevance in practice, especially for high dimensional problems. In contrast to this, e.g., the new SRK scheme RI6 needs 2 evaluations of the drift function  $a$  and only 5 evaluations the diffusion function  $b^j$ ,  $j = 1, \dots, m$ , for each step due to  $s = 3$  stages and  $H_1^{(k)} = \hat{H}_1^{(k)}$ . Further, only  $2m - 1$  independent random variables have to be simulated for each step. The computational effort of the new SRK schemes is even similar to the one of the extrapolated Euler-Maruyama scheme where 3 evaluations of the drift  $a$  and 3 evaluations of the diffusion functions  $b^j$ ,  $j = 1, \dots, m$ , are necessary, and which needs the simulation of  $2m$  independent random variables at each step. However, there are situations where extrapolation methods, especially based on the Euler-Maruyama scheme, have only limited value [9]. This is the case e.g. for stiff problems due to the restricted stability regions of the Euler-Maruyama method [6]. Thus, as in the deterministic setting, higher order one-step methods like the introduced SRK methods are of independent relevance.

As a result of this, the introduced class of SRK methods is of considerable importance, now also for high dimensional problems like e.g. in mathematical finance or physics. Future research may be done by developing efficient SRK schemes of some higher order than order two. Further, a stability analysis and the investigation of implicit SRK methods is of particular interest. Similar to the deterministic setting, embedded SRK schemes can be easily implemented which can be applied with a step size control algorithm, see also [10].

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